

## Differences in COVID-19 vaccination uptake in the first 12 months of vaccine availability in Switzerland – a prospective cohort study

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### Summary

**BACKGROUND:** Widespread vaccination uptake has been shown to be crucial in controlling the COVID-19 pandemic and its consequences on healthcare infrastructures. Infection numbers, hospitalisation rates and mortality can be mitigated if large parts of the population are being vaccinated. However, one year after the introduction of COVID-19 vaccines, a substantial share of the Swiss population still refrains from being vaccinated.

**OBJECTIVES:** We analysed COVID-19 vaccination uptake during the first 12 months of vaccine availability. We compared vaccination rates of different socioeconomic subgroups (e.g., education, income, migration background) and regions (urban vs rural, language region) and investigated associations between uptake and individual traits such as health literacy, adherence to COVID-19 prevention measures and trust in government or science.

**METHODS:** Our analysis was based on self-reported vaccination uptake of a longitudinal online panel of Swiss adults aged 18 to 79 (the “COVID-19 Social Monitor”, analysis sample  $n = 2448$ ). The panel is representative for Switzerland with regard to age, gender, and language regions. Participants have been periodically surveyed about various public health issues from 30 March 2020, to 16 December 2021. We report uptake rates and age-stratified hazard ratios (HRs) by population subgroups without and with additional covariate adjustment using Cox regression survival analysis.

**RESULTS:** Higher uptake rates were found for individuals with more than just compulsory schooling (secondary: unadjusted HR 1.39, 95% confidence interval [CI] 1.10–1.76; tertiary: HR 1.94, 95% CI 1.52–2.47), household income above CHF 4999 (5000–9999: unadj. HR 1.42, 95% CI 1.25–1.61;  $\geq 10,000$  HR 1.99, 95% CI 1.72–2.30), those suffering from a chronic condition (unadj. HR 1.38, 95% CI 1.25–1.53), and for individuals with a sufficient or excellent level of health literacy (sufficient: unadj. HR 1.13, 95% CI 0.98–1.29; excellent: HR 1.21, 95% CI 1.10–1.34). We found lower rates for residents of rural regions (unadj. HR

0.79, 95% CI 0.70–0.88), those showing less adherence to COVID-19 prevention measures, and those with less trust in government or science.

**CONCLUSIONS:** Vaccination uptake is multifactorial and influenced by sociodemographic status, health literacy, trust in institutions and expected risk of severe COVID-19 illness. Fears of unwanted vaccine effects and doubts regarding vaccine effectiveness appear to drive uptake hesitancy and demand special attention in future vaccination campaigns.

### Introduction

SARS-CoV-2 infection numbers and associated morbidity and mortality are still on the rise worldwide. However, the rapid development of effective vaccines against COVID-19 has been a fundamental step towards controlling the pandemic and protecting health systems. Infections, hospitalisations and mortality rates have been shown to be mitigated by vaccinating large parts of the population [1]. Thus, widespread uptake of vaccines is key [2]. However, uptake progressed differently between and within countries owing to challenges in production and distribution of the vaccine, affordable pricing, global allocation and administration of doses [3]. Early in 2021, Israel had administered the most vaccines against COVID-19 per capita worldwide [4]. Also, the UK as well as Canada conducted fast vaccination campaigns [5]. In Europe, by 30 November 2021, Iceland and Portugal show the highest and the Western Balkan countries the lowest vaccination rates. Switzerland, however, shows the lowest vaccination rates compared with other countries in Western and Central Europe. These low rates are similar to the ones found in the United States [6].

### Vaccine hesitancy

After vaccines had become widely available, many countries faced the issue of widespread vaccine hesitancy. Vaccine hesitancy delays immunisation of the population and leads to substantial parts of the population not being vaccinated. Consequently, the benefits of vaccination in terms

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of containing the spread of the pandemic, protecting vulnerable groups and mitigating severe outcomes cannot be fully realised and future outbreaks of COVID-19 might be amplified [7, 8]. The most frequently identified reasons related to vaccine hesitancy include the desire to wait and see if the vaccine is safe, fear of side effects, belief in the superiority of natural immunity, and lack of trust in government and health authorities [9–11]. Sociodemographic differences are important determinants of vaccine acceptance [12, 13]. Systematic reviews found that being female, younger, of lower income and having a lower education level, belonging to an ethnic minority group and living in a rural area were consistently associated with lower intentions to get vaccinated [14, 15].

### The Swiss context

Compared with Israel or the UK, Switzerland had a rather late and slow start to the vaccination campaign. The canton of Lucerne was one of the first to administer vaccines, on 23 December 2020 [16]; the other cantons followed shortly. Priority was given to the elderly and the chronically ill, and, secondly, to healthcare workers and those living with people at risk [17]. By about May 2021, access to vaccines was opened to the general population aged 16 and over, and by the beginning of June 2021, children and adolescents aged 12 to 15 could get vaccinated. Different national and cantonal campaigns were set in place to inform people about the vaccine and to motivate the public and certain subgroups such as the young or foreigners to get vaccinated. From September 2021 on, a “COVID certificate” similar to the EU-issued “Green Pass” for the vaccinated, recovered or negatively tested was declared mandatory to access indoor hospitality venues, cultural, sporting and leisure activities indoors, and large-scale outdoor events [18], putting more pressure on the unvaccinated. Other measures were a national vaccination week at the beginning of November consisting of around 170 additional mobile vaccination centres and personalised advisory services [19]. By 20 December, 69% of the total Swiss population (including children) had received one dose of vaccine according to official statistics, and 67% had been fully vaccinated (generally, two doses; those recovered from COVID-19, one dose) [20]. This places Switzerland among the countries with lower vaccination rates in Europe.

### Aims of this study

Previous Swiss studies focused on vaccination attitudes and willingness to vaccinate, but not on vaccination uptake itself [21–23]. The aim of this study was to investigate the detailed development of COVID-19 vaccination uptake in Switzerland and subgroup heterogeneities over time. The results give detailed insights into socioeconomic and other differences in uptake. Findings contribute to the understanding of the challenges related to widespread vaccine hesitancy and may provide a valuable basis to inform future COVID-19 vaccination or other health prevention campaigns.

## Data and methods

### Design and data

We used data from the COVID-19 Social Monitor, a large cohort study of the Swiss resident population aged 18 to 79 years with online access [24, 25]. The study covers various public health issues and started surveying the population at the beginning of the COVID-19 pandemic. The questionnaire used includes mostly validated items from established population surveys, mainly the Swiss Health Survey (SHS, <https://www.bfs.admin.ch/bfs/de/home/statistiken/gesundheit/erhebungen/SHS.html>), the Swiss Household Panel (SHP, <https://forscenter.ch/projekte/swiss-household-panel>), and the Study on Health, Ageing and Retirement (SHARE, <http://www.share-project.org>). Some items were adapted to fit the context of the pandemic. Questions regarding the COVID-19 pandemic and COVID-19 vaccination uptake were newly developed, expert-reviewed and closely coordinated with another major Swiss COVID-19 population survey, the Corona Immunitas Digital Follow-Up eCohort [26].

Participants were randomly sampled from an online panel whose members have been actively recruited using random probability sampling based on national landline telephone directories and random digit dialing of mobile phone numbers. The sampling process was stratified by age, gender and language region, and is representative of the Swiss resident population (census of 2018) in this regard. An initial sample of 2026 respondents participated in the survey, beginning in March 2020, and was complemented with an additional sample of 1355 from December 2020 onwards. Respondents were surveyed approximately every five to eight weeks during the study period. The last responses included in this analysis were from survey wave 20 with a data-collection period from 6 to 16 December 2021. Table 1 shows the sample size and non-participation rates of the different survey waves. We included 2553 study participants with at least one response to the item on vaccination uptake we introduced on 7 June 2021. We excluded 85 cases (3.3%) because of one or more missing covariates (1 case with missing migration status, 7 for health literacy, 76 for trust in government and trust in science, 2 for trust in science only). Our final analysis sample consists of 2448 cases.

The survey was pseudonymised. The identity of the participants is known only to the panel provider and strictly separated from data collection and the researchers. De-anonymisation of the data was prevented by legal restrictions. Further details on the study methodology and design, as well as baseline characteristics of the sample, are presented in a previously published paper [24]. The data are available under <https://doi.org/10.48620/22>. Vaccination dates were removed from the published data to protect respondents' privacy. We provide the rounded time (in days) from study entry to vaccination, upon request. The research question analysed in this paper had not been anticipated when planning the study and no protocol has been registered.

### Outcome variable

COVID-19 vaccine uptake was elicited from wave 17 onwards (7 June 2021) by asking respondents retrospectively whether and when they received their first vaccine dose (“Have you been vaccinated against the coronavirus (at least one dose)?”, “When have you received your first dose?”). First-dose vaccination is a valid measure of vaccine uptake as there is so far no evidence of any non-uptake of second doses on a relevant scale. Further, this operationalisation also fits for those with a prior SARS-CoV-2 infection who were recommended to get only one dose of vaccine, and to those few receiving the Johnson and Johnson vaccine that requires only one dose. Respondents who had not received a first dose of the vaccine when last surveyed were asked about the reasons for not being vaccinated.

### Sociodemographic subgroups and predictors

For subgroup comparisons and as potential determinants for uptake, we included three sets of predictors, which we selected a priori based on theoretical considerations. Besides basic sociodemographics we included health literacy, adherence to COVID-19 prevention measures, and trust, which are all highlighted as important determinants of COVID-19 vaccine uptake in the literature:

1. Basic sociodemographics: age (18 to 29; 30 to 39; 40 to 49; 50 to 59; 60 to 79), gender (male; female), completed education (compulsory schooling; secondary degree II; tertiary degree), household income (CHF <5000; CHF 5000–9999; CHF ≥10,000; no answer), place of residence (urban; rural), language region (German-, French-, Italian-speaking), and migration background (Swiss-born; foreign-born), chronic condition (yes; no) including hypertension, cardiovascular diseases, diabetes, respiratory diseases, cancer and renal disease, excluding mental health disorders and arthrosis/arthritis.

2. Health literacy (problematic/inadequate; sufficient; excellent) measured with an adapted version of the HLS-EU-

Q12 scale [27, 28] and, as an indicator for adherence to COVID-19 prevention measures, whether respondents refrained from domestic visits to friends and relatives to protect themselves and others from the coronavirus (rarely; sometimes; mostly). We used responses elicited in December 2020, a time with high infection rates and partial lockdown, when vaccines were not yet available in Switzerland. For respondents not participating in that survey wave, we used responses from previous waves (210 cases).

3. Trust in government and trust in science (“Regarding the COVID-19 pandemic, how much do you trust the following sources?” Swiss government (e.g. Federal Office of Public Health); Science) with the categories low (“little”/“very little”), middle (“moderately”), and high (“strongly”/“very strongly”). We used responses elicited in December 2020, just before the start of the vaccination campaign. For respondents not participating in that survey wave, we used responses from previous waves, or, if none were available, from subsequent follow-up waves (133 cases for trust in government, 134 cases for trust in science).

### Statistical analysis

We calculated COVID-19 vaccine uptake rates as the proportion of participants having received a first dose at a particular point in time up to the end of the study period on 16 December 2021 using Kaplan-Meier estimation. Results are reported as Kaplan-Meier curves and as rates by the end of the study period.

We then compared uptake rates over time of the above-specified subgroups using Cox regression survival analysis. Cox regression allows for a relative comparison of uptake rates at different points in time and accounts for right-censoring. The observation period under consideration started on 19 December 2020, the day when vaccines first became available in Switzerland. The observation period ended with an individual’s last survey participation, i.e., at the latest with the end of the study period on 16 December 2021, or the date of vaccination, whichever came

**Table 1:**  
Overview of COVID-19 Social Monitor sample size and nonparticipation by wave.

Survey wave	1	2	3	4	5	6	7
Survey start	30 March 2020	6 April 2020	14 April 2020	27 April 2020	11 May 2020	25 May 2020	15 June 2020
No. of initial sample	2026	2026	2026	2026	2026	2026	2026
No. of additional sample	0	0	0	0	0	0	0
No. of combined sample	2026	2026	2026	2026	2026	2026	2026
No. of participants	2026	1537	1540	1729	1673	1616	1522
Nonparticipation (%)	0	24	24	15	17	20	25
Survey wave	8	9	10	11	12	13	14
Survey start	14 July 2020	17 August 2020	28 September 2020	9 November 2020	14 December 2020	25 January 2021	22 February 2021
No. of initial sample	2026	2026	2026	2026	2026	2026	2026
No. of additional sample	0	0	0	0	1355	1355	1355
No. of combined sample	2026	2026	2026	2026	3381	3381	3381
No. of participants	1508	1532	1511	1492	2802	2564	2346
Nonparticipation (%)	26	24	25	26	17	24	31
Survey wave	15	16	17	18	19	20	
Survey start	29 March 2021	3 May 2021	7 June 2021	30 August 2021	18 October 2021	6 December 2021	
No. of initial sample	2026	2026	2026	2026	2026	2026	
No. of additional sample	1355	1355	1355	1355	1355	1355	
No. of combined sample	3381	3381	3381	3381	3381	3381	
No. of participants	2219	2154	2095	1921	1947	1951	
Nonparticipation (%)	34	36	38	43	42	42	

first. To account for the age-specific availability of vaccines and vaccination progress, we stratified all analyses by age (categorised by 18 to 29, 30 to 39, 40 to 49, 50 to 59, and 60 to 79 years). This allows for different baseline hazard ratios (HRs) by age group, while keeping HRs for other predictors equal. We estimated three multivariable models using the described predictor sets. We report unadjusted (from univariable regression models) HRs and adjusted (from multivariable) HRs with 95% confidence intervals (CIs). We tested the proportional-hazards assumption using Schoenfeld's test. We found a violation of the assumption for gender ( $p = 0.02$ ), the Italian vs the German language region ( $p = 0.04$ ), and for excellent vs problematic/inadequate health literacy ( $p = 0.04$ ); the global  $p$ -value was  $<0.001$ . Consequently, we carried out separate analyses by gender to check the robustness of our results (see supplementary tables S1 and S2 in the appendix).

Finally, we report the reasons for not being vaccinated as proportions of the nonvaccinated according to responses in the last survey wave (overall, and by gender). Data analysis was carried out with Stata/SE version 17.0, the analysis do-files are provided in the supplement S4.

We used weights to correct for sampling and attrition bias as described in Moser et al. [25]. In brief, sampling weights were constructed using variables age, gender, and language region and the 2018 Swiss census data as reference. Attrition weights used additional information about employment status, living with a partner or not, and highest attained education. Details on the weighting strategy including a sensitivity analysis regarding alternative modelling strategies are provided in supplement S3.

### Ethics approval and consent

The Cantonal Ethics Commission of Zurich concluded that the current study does not fall within the scope of the Human Research Act (BASEC-Nr. Req-2020-00323). The study is a pseudonymised survey. Participants gave their general consent to be part of research studies when accepting the panel provider's invitation to the online panel from which we sampled respondents. Explicit informed consent was therefore not needed from participants for this study. Participation in the study was voluntary and participants could always withdraw.

## Results

### Study population

Table 2 shows the characteristics of study participants and the resulting proportions when applying sampling weights. We analysed responses from 2448 study participants. Half of the study population was 50 years or older (54%) with an equal distribution of men and women. A majority of study participants had secondary school as the highest achieved education level (68%), reported a monthly income of below CHF 10,000 (70%), had no chronic disease condition (65%) and lived in an urban area (80%).

### Vaccination progress

By the end of the study period on 16 December 88% of the study population had received their first dose of

COVID-19 vaccine (fig. 1). Vaccinations had started by the end of December 2020; rates progressed slowly until April and then increased steeply until the end of June when progress (first dose) slowed down considerably after having reached an overall prevalence of about 70% of the study population. The slowdown began earlier and was particularly marked among the elderly who had early access to vaccination, whereas rates for the younger with delayed eligibility continued to soar until about October. After October, rates continued to increase constantly but only slowly until the end of the study period.

Vaccination uptake differed considerably between age groups; in particular, the elderly aged 60 to 79 years showed a much earlier and faster uptake. Between other soci-demographic groups there were differences too, albeit on a smaller scale: respondents with compulsory schooling, with lower household income as well as those living in rural regions showed a slower uptake and a lower rate at the end of the study period. Respondents with compulsory schooling showed the lowest rate at 81%, those in the income group CHF  $\geq 10,000$  the highest at 94%.

Compared to official data provided by the Federal Office of Public Health [20], our sample showed considerably higher vaccination rates. Depending on age group, rates were between 4 (age group 70 to 79 years) and 15 percentage points (20 to 29) higher than the official vaccination rates (table 3).

### Cox regression results

Table 4 shows uptake rates and HRs from age-stratified univariable and age-stratified multivariable analysis of the three models, including different sets of predictors. Unadjusted HRs from the age-stratified univariable analysis, without additional covariates, showed that uptake was lower for females than males (HR 0.89, 95% CI 0.81–0.97), increased with educational level (secondary: HR 1.39, 95% CI 1.10–1.76; tertiary: HR 1.94, 95% CI 1.52–2.47) and household income (CHF 5000–9999, HR 1.42, 95% CI 1.25–1.61; CHF  $\geq 10,000$ : HR 1.99, 95% CI 1.72–2.30), and was lower for respondents residing in rural vs in urban regions (HR 0.79, 95% CI 0.70–0.88). No differences could be found for migration background and language region. Individuals with a chronic condition showed a higher uptake rate than those without (HR 1.38, 95% CI 1.25–1.53).

A separate analysis by gender (see supplementary tables S1 and S2) showed some few gender-specific differences: women living in rural areas had an even lower vaccination uptake than their male counterparts; men with lower health literacy showed a lower uptake than those with higher health literacy, whereas there was no statistically significant difference for females in this regard. For other predictors, we see no substantial differences.

Results from the multivariable model including basic sociodemographic predictors (model 1) are comparable to the ones from the univariable analysis with some slightly smaller hazard ratios. In model 2, after adjustment for sociodemographics, a higher level of health literacy showed a positive association with uptake (sufficient: HR 1.07, 95% CI 0.93–1.23; excellent: HR 1.19, 95% CI 1.08–1.32; ref-



erence category: problematic/inadequate), as did stronger adherence to COVID-19 prevention measures (refrained from domestic visits). In multivariable model 3, which additionally included trust in government and science, a strong positive association with both trust in government and trust in science was found.

### Reasons against vaccination

The 12% of respondents not being vaccinated according to the last survey wave with responses collected from 6 to 16 December indicated fears of side effects (reported by 57% of the not vaccinated), doubts regarding the effectiveness of the vaccine (57%), and a preference for natural/traditional remedies (36%) as main reasons against the vaccine. Twenty-one percent reported having had a COVID-19 infection as a reason against a vaccination (fig. 2). Only 4% mentioned medical reasons. Separate analysis by gender showed a similar pattern, with females reporting slightly more often fears of side effects and having had a COVID-19 infection and with males questioning more frequently the vaccines' effectiveness and whether COVID-19 posed an actual threat to them ("A COVID-19 infection is not dangerous for me.").

## Discussion

### Main results

Although vaccine uptake in Switzerland progressed quickly after a slow beginning due to a lack of vaccine availability, uptake slowed down by the end of June 2021, when about 70% in our study population had received their first dose. Vaccine availability was no longer an issue at that time, but significant vaccine hesitancy, indifference and, possibly, implicit accessibility barriers among all age groups curbed the further increase in uptake rates. Most apparent differences in uptake could be found between respondents with low and high levels of completed education, low and high household income, rural vs urban regions of residence, and between those with and without a chronic condition. None or only small differences could be found between genders, the language regions and between respondents with and without migration background. Lower vaccine uptake was associated with lower health literacy (for males, but not for females), lower adherence to COVID-19 prevention measures, as well as with low levels of trust in government and science.

**Table 2:**  
Characteristics of study population (n = 2448).

		n	Proportion (in %) *
Age	18 to 29	434	15
	30 to 39	353	12
	40 to 49	468	17
	50 to 59	619	24
	60 to 79	574	30
Gender	Male	1263	50
	Female	1185	50
Education	Compulsory	107	5
	Secondary	1629	68
	Tertiary	712	27
Household income	CHF <5,000	552	23
	CHF 5000–9,999	1139	47
	≥10,000 CHF	494	20
	No answer	263	10
Migration background	Swiss-born	2142	88
	Foreign-born	306	12
Place of residence	Urban	1987	80
	Rural	461	20
Language region	German	1591	72
	French	495	23
	Italian	362	4
Chronic condition	No	1662	65
	Yes	786	35
Health literacy	Problematic/Inadequate	1075	42
	Sufficient	336	15
	Excellent	1037	43
Adherence to prevention measures (refrained from visits)	Rarely	413	17
	Sometimes	418	17
	Mostly	1617	66
Trust in government	Low	276	11
	Middle	523	21
	High	1649	69
Trust in science	Low	236	10
	Middle	617	25
	High	1595	65

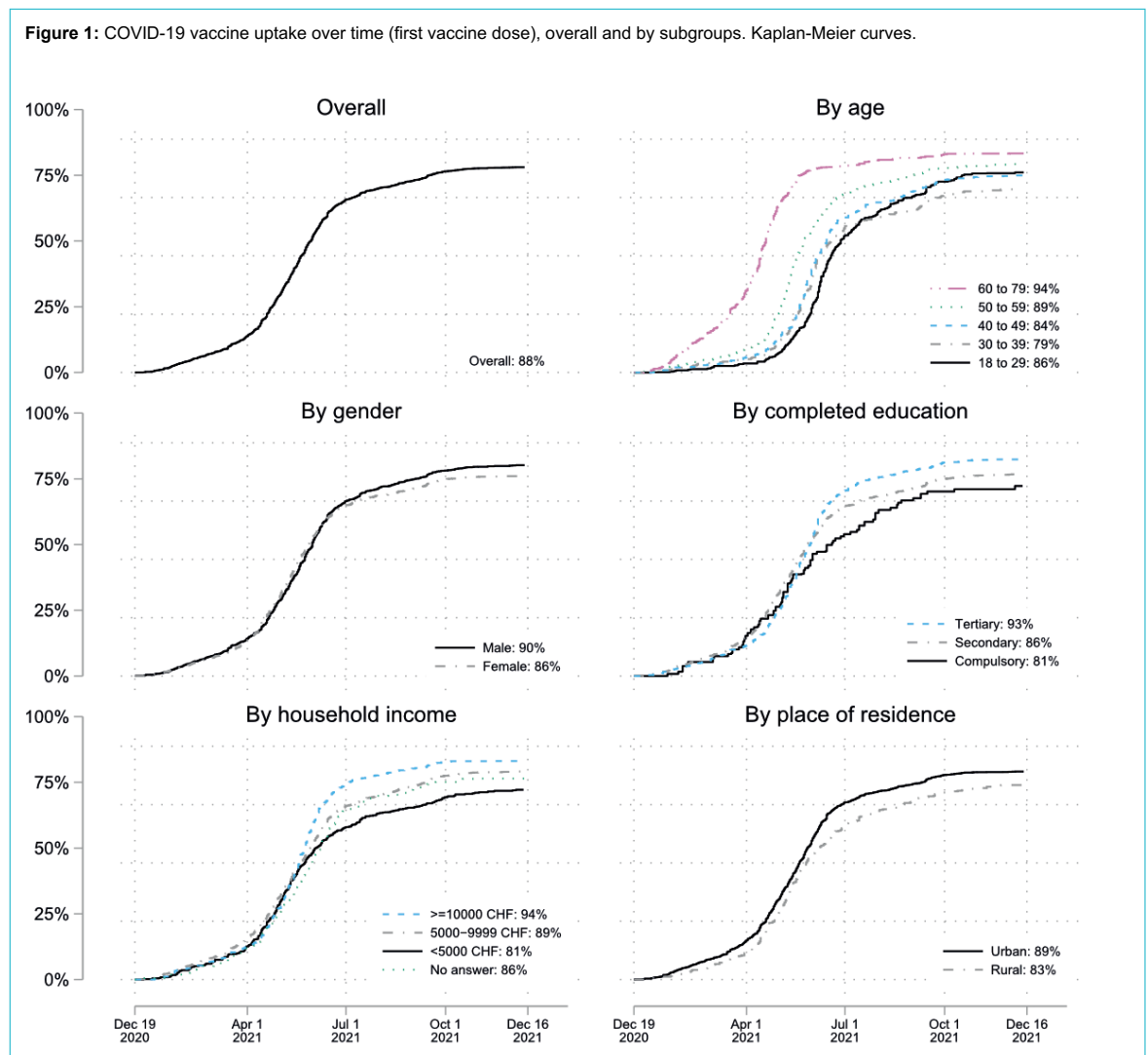
\* Using weighted sample

**Explaining the pattern and relation to previous findings**

Similar patterns regarding age groups, household income, education level, gender and migration background were observed in analyses based on data from a nationwide Swiss seroprevalence study [29]. Moreover, widespread vaccine hesitancy has been observed in Switzerland for other vaccines, such as human papillomavirus (HPV) [7, 30], measles [31] and influenza [32]. These studies found a pattern similar to our results with regard to lower vaccination uptake in rural compared with urban areas [30–32]. For HPV vaccines, fear of side effects and general opposition to vaccination were two main reasons for not being vaccinated [33], but also accessibility to HPV vaccines, i.e. the existence and the scope of cantonal vaccination programmes led to differences in vaccination rates [34, 35]. Accessibility has likely played a role in

COVID-19 vaccination uptake too. (Online) registration procedures, as well as problems with getting time off work to get vaccinated, might have posed something of a hurdle to persons with lower education levels or a migration background, in particular when coupled with a lack of social support or a nonexistent established primary care relationship.

Interestingly, we did not find systematic differences between the language regions, even though pronounced differences between Swiss cantons exist, with the vaccination rate for the total population ranging from 55.9–70.7% [20]. Such differences might be explained by local context and the federal organisation of the vaccination programmes in Switzerland. Neighbouring countries show large differences in vaccination rates: Germany (71.2%) and Austria (70.1%) have lower vaccination rates for the total population compared with Italy (74.3%) and France (73.6%)



**Table 3:**

COVID-19 vaccine uptake of the study population and of the general population according to official vaccination rates provided by the Federal Office of Public Health (FOPH) by age group as of 16 December 2021.

	Age groups (in years)					
	20–29	30–39	40–49	50–59	60–69	70–79
Official rates (FOPH)	71%	73%	77%	81%	86%	91%
Study sample	86%	79%	84%	89%	93%	95%

Rates for study sample are based on weighted Kaplan-Meier estimates.

(see [36]). This, however, does not seem to translate into sizeable differences between the Swiss language regions. Whereas raw uptake rates are almost identical between the language regions, hazard ratios for the Italian language region are <1 (albeit not significantly different from unity at the 5% level). This is a result of a seemingly lower vaccination rate in the Ticino after control for age (age-stratified univariable model) and other covariates (age-stratified multivariable models).

In the case of COVID-19, extensive misinformation since the beginning of the pandemic nourished mistrust in science and public health authorities [37]. Hence, the type of sources trusted and used to obtain vaccination-related information plays a crucial role in decisions about COVID-19 vaccination as the association between misinformation and vaccine hesitancy is well documented

[12, 38]. Furthermore, even though the relationship between health literacy and uptake of childhood vaccines or influenza vaccination remains unclear [39], levels of low health literacy have been shown to be related to COVID-19 vaccine hesitancy [37].

Our data provide no information about the effectiveness of public health measures intended to increase vaccination rates. Presently, using incentives and penalties to encourage vaccination is intensively debated. Research has shown that even small financial incentives can strongly increase the usage of a COVID-19 contact tracing app [40]. Incentives such as remunerations, vaccination possibilities at local doctors' practices or providing more freedom have been shown in a hypothetical factorial survey experiment in Germany to be effective, but to different degrees depending on the age of respondents [41]. Also, manda-

**Table 4:** COVID-19 vaccine uptake and Cox proportional hazard ratios from univariable and multivariable analyses (N=2,448).

	Uptake <sup>a</sup> (%)	Unadj HR <sup>b</sup>	(95% CI)	Adj HR Model 1 <sup>c</sup>	(95% CI)	Adj HR Model 2 <sup>c</sup>	(95% CI)	Adj HR Model 3 <sup>c</sup>	(95% CI)
Gender									
Male	90	1.00		1.00		1.00		1.00	
Female	86	0.89	(0.81–0.97)	0.97	(0.88–1.06)	0.94	(0.86–1.03)	0.92	(0.84–1.01)
Education									
Compulsory	81	1.00		1.00		1.00		1.00	
Secondary	86	1.39	(1.10–1.76)	1.23	(0.98–1.55)	1.23	(0.97–1.56)	1.12	(0.89–1.41)
Tertiary	93	1.94	(1.52–2.47)	1.59	(1.25–2.02)	1.58	(1.24–2.03)	1.32	(1.03–1.68)
Household income									
CHF <5000	81	1.00		1.00		1.00		1.00	
CHF 5000–9999	89	1.42	(1.25–1.61)	1.37	(1.21–1.56)	1.35	(1.19–1.54)	1.28	(1.13–1.46)
CHF ≥10,000	94	1.99	(1.72–2.30)	1.81	(1.55–2.10)	1.76	(1.51–2.05)	1.69	(1.45–1.98)
No answer	86	1.39	(1.16–1.66)	1.40	(1.17–1.67)	1.38	(1.15–1.65)	1.36	(1.14–1.63)
Migration background									
Swiss-born	88	1.00		1.00		1.00		1.00	
Foreign-born	88	1.04	(0.90–1.20)	0.99	(0.86–1.14)	0.97	(0.84–1.12)	0.95	(0.83–1.10)
Place of residence									
Urban	89	1.00		1.00		1.00		1.00	
Rural	83	0.79	(0.70–0.88)	0.80	(0.72–0.90)	0.79	(0.70–0.88)	0.78	(0.70–0.88)
Language region									
German	88	1.00		1.00		1.00		1.00	
French	89	1.04	(0.93–1.16)	1.00	(0.89–1.12)	1.03	(0.92–1.15)	0.99	(0.88–1.11)
Italian	89	0.94	(0.84–1.05)	0.90	(0.79–1.01)	0.88	(0.78–1.00)	0.89	(0.79–1.00)
Chronic condition									
No	85	1.00		1.00		1.00		1.00	
Yes	94	1.38	(1.25–1.53)	1.41	(1.27–1.56)	1.39	(1.25–1.53)	1.38	(1.25–1.53)
Health literacy									
Problematic/inadequate	85	1.00				1.00		1.00	
Sufficient	91	1.13	(0.98–1.29)			1.07	(0.93–1.23)	1.01	(0.88–1.16)
Excellent	90	1.21	(1.10–1.34)			1.19	(1.08–1.32)	1.10	(0.99–1.21)
Adherence to prevention measures (refrained from visits)									
Rarely	75	1.00				1.00		1.00	
Sometimes	88	1.56	(1.33–1.84)			1.56	(1.32–1.85)	1.36	(1.15–1.61)
Mostly	91	1.67	(1.46–1.91)			1.67	(1.46–1.91)	1.42	(1.24–1.62)
Trust in government									
Low	63	1.00						1.00	
Middle	81	1.76	(1.45–2.14)					1.41	(1.14–1.74)
High	94	2.70	(2.28–3.20)					1.74	(1.42–2.14)
Trust in science									
Low	65	1.00						1.00	
Middle	81	1.69	(1.39–2.05)					1.32	(1.06–1.65)
High	94	2.65	(2.21–3.17)					1.64	(1.31–2.05)

CI confidence interval; HR hazard ratio.

<sup>a</sup> Uptake based on weighted Kaplan-Meier estimates as of last common censoring date across all subgroups (8 December); <sup>b</sup> stratified by age; <sup>c</sup> stratified by age and adjusted for all the other variables included in the model as shown.

tory COVID-19 certificates seem to increase vaccination uptake, especially among the young [42]. However, the overall effect of such measures is highly context-specific and hard to predict for the current Swiss situation.

### Implications

The fact that individuals with lower educational levels and lower health literacy show a lower COVID-19 vaccination uptake, and that fears of side effects and doubts about the efficacy of the vaccine are prominent reasons against vaccination suggests that the risk and benefits of being vaccinated should be communicated more broadly and in a more comprehensible way. Also, vaccination campaigns require a determined effort from local/cantonal authorities and should be planned from early on to be as accessible as possible for the whole population, including persons with poor local language and/or reading skills, limited time availability, or in otherwise disadvantaged situations that might pose barriers to vaccination [43]. As lower uptake is associated with low trust in government and science, actors other than government agencies and scientists might be more successful in communicating the benefits of being vaccinated to subgroups with low uptake rates. Local doctors and local politicians or other recognised public persons might be better suited. Implementing strategies based on behavioural science insights such as nudge techniques or giving people a choice (between various vaccines) might be options to consider [44].

### Strengths and limitations

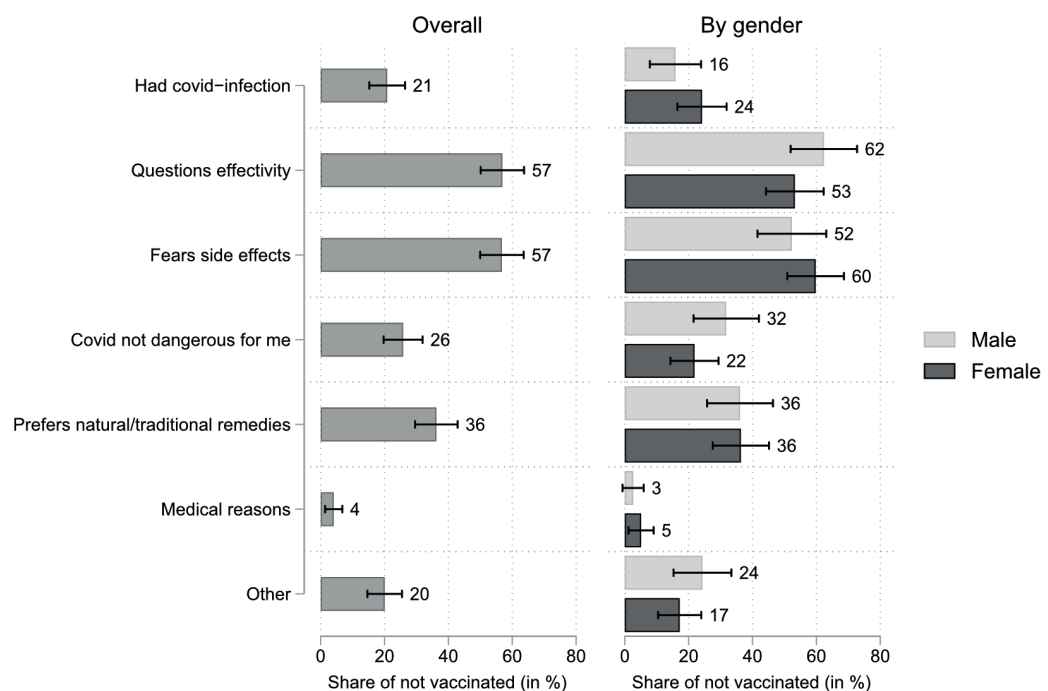
A strength of this study is that not much pressure was put on the unvaccinated in Switzerland; hence, our data reflect well the study populations' preferences and attitudes to-

wards vaccination. Of course, our study has several limitations such as the selectivity of the sample, which is restricted to persons with online access, good knowledge of Italian, French or German, and capabilities as well as willingness to participate in an online survey on health and other topics related to the pandemic. Persons with migration experience are underrepresented. This selectivity might explain our relatively higher overall vaccination rate compared with official data. Selectivity might have increased with study duration, due to non-random dropouts. Also, self-reporting of the vaccination status might be prone to social desirability bias, even though this might be less of a problem in an online survey and for a behaviour far from being marginal in the surveyed population [45].

### Conclusions

After one year of vaccine availability, the COVID-19 vaccine uptake in Switzerland reached a plateau in Switzerland. Uptake rates differed considerably between population subgroups with the less educated, less well-off, and rural population showing the lowest rates. To increase uptake rates and to strengthen future vaccination booster campaigns, efforts should be continued to mitigate fears of unwanted vaccine effects and doubts regarding vaccine effectiveness, and to point out the individual and societal benefits of vaccination – also for healthy, younger individuals. Specifically, targeted measures communicating the risk and benefits of being vaccinated relative to not being vaccinated in a comprehensible way and by trusted individuals might help to counteract low vaccination rates among these subgroups.

**Figure 2:** Stated reasons for not being vaccinated at the end of study period overall and by gender. Multiple responses possible. n = 227 (94 males, 133 females), cases weighted to correct for sampling and wave-nonresponse.





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**Potential competing interests**

All authors have completed and submitted the International Committee of Medical Journal Editors form for disclosure of potential conflicts of interest. No potential conflict of interest was disclosed.

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## Appendix: Supplementary material

Table S1:

Males – COVID-19 vaccine uptake and Cox proportional hazard ratios from univariable and multivariable analyses (n = 1263).

	Uptake <sup>a</sup> (%)	Unadj HR <sup>b</sup>	(95% CI)	Adj HR Model 1 <sup>c</sup>	(95% CI)	Adj HR Model 2 <sup>c</sup>	(95% CI)	Adj HR Model 3 <sup>c</sup>	(95% CI)
Education									
Compulsory	81	1.00		1.00		1.00		1.00	
Secondary	89	1.49	(1.05–2.11)	1.44	(1.00–2.07)	1.44	(0.99–2.09)	1.32	(0.91–1.92)
Tertiary	94	2.16	(1.51–3.10)	1.98	(1.37–2.87)	1.97	(1.34–2.90)	1.66	(1.13–2.44)
Household income									
CHF <5000	87	1.00		1.00		1.00		1.00	
CHF 5000–9999	90	1.23	(1.04–1.45)	1.18	(1.00–1.40)	1.16	(0.97–1.38)	1.13	(0.94–1.35)
CHF ≥10,000	94	1.75	(1.45–2.11)	1.56	(1.28–1.90)	1.53	(1.25–1.87)	1.54	(1.25–1.90)
No answer	91	1.49	(1.11–1.98)	1.52	(1.15–2.02)	1.47	(1.12–1.95)	1.53	(1.17–2.01)
Migration background									
Swiss-born	90	1.00		1.00		1.00		1.00	
Foreign-born	92	1.11	(0.93–1.32)	1.09	(0.91–1.31)	1.05	(0.88–1.26)	1.00	(0.83–1.20)
Place of residence									
Urban	90	1.00		1.00		1.00		1.00	
Rural	90	0.88	(0.76–1.02)	0.87	(0.75–1.01)	0.84	(0.72–0.98)	0.83	(0.71–0.96)
Language region									
German	91	1.00		1.00		1.00		1.00	
French	89	1.03	(0.88–1.21)	0.98	(0.83–1.15)	1.01	(0.86–1.19)	0.96	(0.81–1.13)
Italian	90	0.88	(0.74–1.03)	0.80	(0.66–0.96)	0.80	(0.67–0.96)	0.79	(0.67–0.94)
Chronic condition									
No	88	1.00		1.00		1.00		1.00	
Yes	94	1.32	(1.15–1.52)	1.41	(1.22–1.62)	1.36	(1.19–1.57)	1.35	(1.18–1.55)
Health literacy									
Problematic/inadequate	86	1.00				1.00		1.00	
Sufficient	96	1.32	(1.10–1.59)			1.25	(1.04–1.51)	1.17	(0.97–1.42)
Excellent	93	1.38	(1.21–1.58)			1.31	(1.14–1.50)	1.19	(1.05–1.36)
Adherence to prevention measures (refrained from visits)									
Rarely	77	1.00				1.00		1.00	
Sometimes	90	1.62	(1.29–2.04)			1.63	(1.29–2.06)	1.48	(1.17–1.85)
Mostly	94	1.72	(1.43–2.05)			1.71	(1.42–2.05)	1.48	(1.23–1.79)
Trust in government									
Low	70	1.00						1.00	
Middle	84	1.68	(1.31–2.16)					1.38	(1.05–1.83)
High	96	2.48	(2.01–3.06)					1.60	(1.22–2.09)
Trust in science									
Low	71	1.00						1.00	
Middle	86	1.61	(1.25–2.08)					1.30	(0.96–1.76)
High	95	2.44	(1.93–3.09)					1.58	(1.16–2.16)

CI: confidence interval; HR: hazard ratio.

<sup>a</sup> Uptake based on weighted Kaplan-Meier estimates as of last common censoring date across all subgroups (8 December)<sup>b</sup> Stratified by age<sup>c</sup> Stratified by age and adjusted for all the other variables included in the model as shown.

**Table S2:**

Females – COVID-19 vaccine uptake and Cox proportional hazard ratios from univariable and multivariable analyses (N=1,185).

	Uptake <sup>a</sup> (%)	Unadj HR <sup>b</sup>	(95% CI)	Adj HR Model 1 <sup>c</sup>	(95% CI)	Adj HR Model 2 <sup>c</sup>	(95% CI)	Adj HR Model 3 <sup>c</sup>	(95% CI)
<b>Education</b>									
Compulsory	83	1.00		1.00		1.00		1.00	
Secondary	84	1.31	(0.96–1.79)	1.12	(0.83–1.52)	1.13	(0.83–1.53)	1.01	(0.74–1.39)
Tertiary	91	1.70	(1.22–2.38)	1.35	(0.97–1.88)	1.34	(0.96–1.86)	1.10	(0.78–1.54)
<b>Household income</b>									
CHF <5000	77	1.00		1.00		1.00		1.00	
CHF 5000–9999	88	1.56	(1.30–1.87)	1.57	(1.31–1.89)	1.55	(1.30–1.86)	1.43	(1.19–1.72)
CHF ≥10,000	94	2.19	(1.76–2.72)	2.06	(1.63–2.61)	2.00	(1.58–2.52)	1.79	(1.41–2.28)
No answer	83	1.39	(1.10–1.75)	1.38	(1.10–1.74)	1.38	(1.10–1.74)	1.30	(1.03–1.65)
<b>Migration background</b>									
Swiss-born	86	1.00		1.00		1.00		1.00	
Foreign-born	84	0.97	(0.77–1.21)	0.90	(0.71–1.13)	0.88	(0.70–1.12)	0.92	(0.74–1.16)
<b>Place of residence</b>									
Urban	88	1.00		1.00		1.00		1.00	
Rural	78	0.73	(0.61–0.86)	0.75	(0.63–0.89)	0.74	(0.62–0.88)	0.74	(0.62–0.88)
<b>Language region</b>									
German	85	1.00		1.00		1.00		1.00	
French	88	1.06	(0.90–1.24)	1.01	(0.86–1.18)	1.03	(0.88–1.20)	1.01	(0.86–1.18)
Italian	89	1.00	(0.86–1.18)	1.00	(0.85–1.18)	0.98	(0.83–1.15)	1.00	(0.85–1.18)
<b>Chronic condition</b>									
No	82	1.00		1.00		1.00		1.00	
Yes	93	1.42	(1.22–1.65)	1.41	(1.22–1.64)	1.41	(1.22–1.63)	1.43	(1.23–1.65)
<b>Health literacy</b>									
Problematic/Inadequate	84	1.00				1.00		1.00	
Sufficient	87	1.01	(0.83–1.23)			0.94	(0.77–1.16)	0.91	(0.74–1.11)
Excellent	87	1.10	(0.95–1.27)			1.08	(0.94–1.25)	1.02	(0.88–1.18)
<b>Adherence to prevention measures (refrained from visits)</b>									
Rarely	73	1.00				1.00		1.00	
Sometimes	86	1.53	(1.20–1.94)			1.47	(1.16–1.87)	1.23	(0.97–1.57)
Mostly	89	1.62	(1.33–1.99)			1.58	(1.30–1.93)	1.31	(1.07–1.60)
<b>Trust in government</b>									
Low	53	1.00						1.00	
Middle	78	2.05	(1.50–2.82)					1.60	(1.13–2.28)
High	92	3.27	(2.45–4.37)					2.04	(1.46–2.86)
<b>Trust in science</b>									
Low	57	1.00						1.00	
Middle	78	1.86	(1.38–2.53)					1.37	(0.98–1.92)
High	93	2.99	(2.26–3.96)					1.73	(1.25–2.40)

CI: confidence interval; HR: hazard ratio.

<sup>a</sup> Uptake based on weighted Kaplan-Meier estimates as of last common censoring date across all subgroups (8 December)<sup>b</sup> Stratified by age<sup>c</sup> Stratified by age and adjusted for all the other variables included in the model as shown.

### S3. Statistical methods for calibration weights.

The following section has been already described in the Supplementary Material (S2 Text: Statistical methods for calibration weights) in Moser et al. (<https://doi.org/10.1371/journal.pone.0256253>).

We describe the survey population and calibrations weights by frequencies (n), percentages (%) and box plots. We calculate 95% confidence intervals (CI) for proportions based on the method from Clopper and Pearson. Let  $\mathcal{O}_{1355} = (\mathcal{O}_{1,1355}, \dots, \mathcal{O}_{1355,1355})^T$  be the set of observed outcomes from the additional December 2020 survey sample of 1,355 participants and  $\mathcal{O}_{2026} = (\mathcal{O}_{1,2026}, \dots, \mathcal{O}_{2026,2026})^T$  be the set of observed outcomes from the March 2020 survey sample. Because individuals were randomly selected we assume that complete set of observed variables  $\mathcal{O}_{3381} = \{\mathcal{O}_{1355}, \mathcal{O}_{2026}\}$  are realisation from identically and independently distributed random variables. We denote the combined initial March 2020 survey sample of 2,026 participants and the additional 1,355 participants from the December 2020 survey sample as December\* 2020 sample. Note that - because of nonresponse of participants from the March 2020 survey - the December\* 2020 survey sample is actually not fully observed.

Let  $Y_{il,u}$  denote the binary outcome whether individual  $i \leq n_u$  living in language region  $l = \{1, 2, 3\}$  was sampled from the Swiss population aged 15 years or older in 2018 in survey wave  $u \in \{1 = \text{March}, 2 = \text{December}\}$ . We construct sampling weights  $SW_{i,u} := SW_{il,u} = 1/Prob(Y_{il,u} = 1)$  for  $i \leq n_u, l = \{1, 2, 3\}, u \in \{1, 2\}$ . We specify a hierarchical logistic regression model (M1) as

$$(M1) \quad Y_{il,u} = \beta_0 + \beta^T \text{agecat} * \text{gender} + lregion_l + \epsilon_{il}, \quad i \leq n_u, l = \{1, 2, 3\}, u \in \{1, 2\},$$

where  $n_u$  is the survey sample size in  $u$ ,  $\beta_0$  is the overall intercept,  $\beta^T = (\beta_1, \beta_2, \dots, \beta_5)$  is a vector of estimates from an interaction term (a+b+a:b) between age category (*agecat1*: <45 years, *agecat2*: 45- <65 years, *agecat3*: 65+ years) and gender (*female*: 0 Men, 1 Women), and  $lregion_l \sim N(0, \tau_{lregion})$  is an unstructured random effect for the variable language region (*lregion*: 1 German/Romansh, 2 French, 3 Italian). We use a Bayesian modeling approach using noninformative centered Gaussian priors for  $\beta_0, \beta^T$ , i.e.  $p(\beta_z | \mu_z, \sigma_z^2) = N(0, 1000)$ ,  $z \leq 5$ , and a centered Gaussian distributed prior with an inverse-Gamma distributed log precision parameter  $\tau_{lregion} \sim \Gamma^{-1}(1, 5/100000)$  for the unstructured random effect.

As sensitivity analyses we perform 1) a Bayesian hierarchical logistic model (M2) with age category, gender and language region as random effects, 2) a Bayesian non-hierarchical logistic regression model (M3) with age category, gender and language region as fixed effects, 3) a frequentist logistic regression model (see the appendix section sensitivity analysis in this document). We compare performances of models (M1), (M2) and (M3) by Watanabe–Akaike information criterion (WAIC) with a lower WAIC indicating better model performance.

Let  $Z_{iw,u} \in \{0, 1\}$  denote the binary outcome whether individual  $i \leq n_u$  participated in survey wave  $w \geq 2$ ,  $u \in \{1, 2\}$ . We define a Bayesian logistic regression model for the construction of nonresponse weights as follows

$$Z_{iw,u} = \gamma_0 + \gamma_1 \text{agecat} + \gamma_2 \text{gender} + \gamma_3 \text{lregion} + \gamma_4 \text{partner} + \gamma_5 \text{work} + \gamma_6 \text{education} + \epsilon_{iw}, \quad i \leq n_u, w \geq 2, u \in \{1, 2\},$$

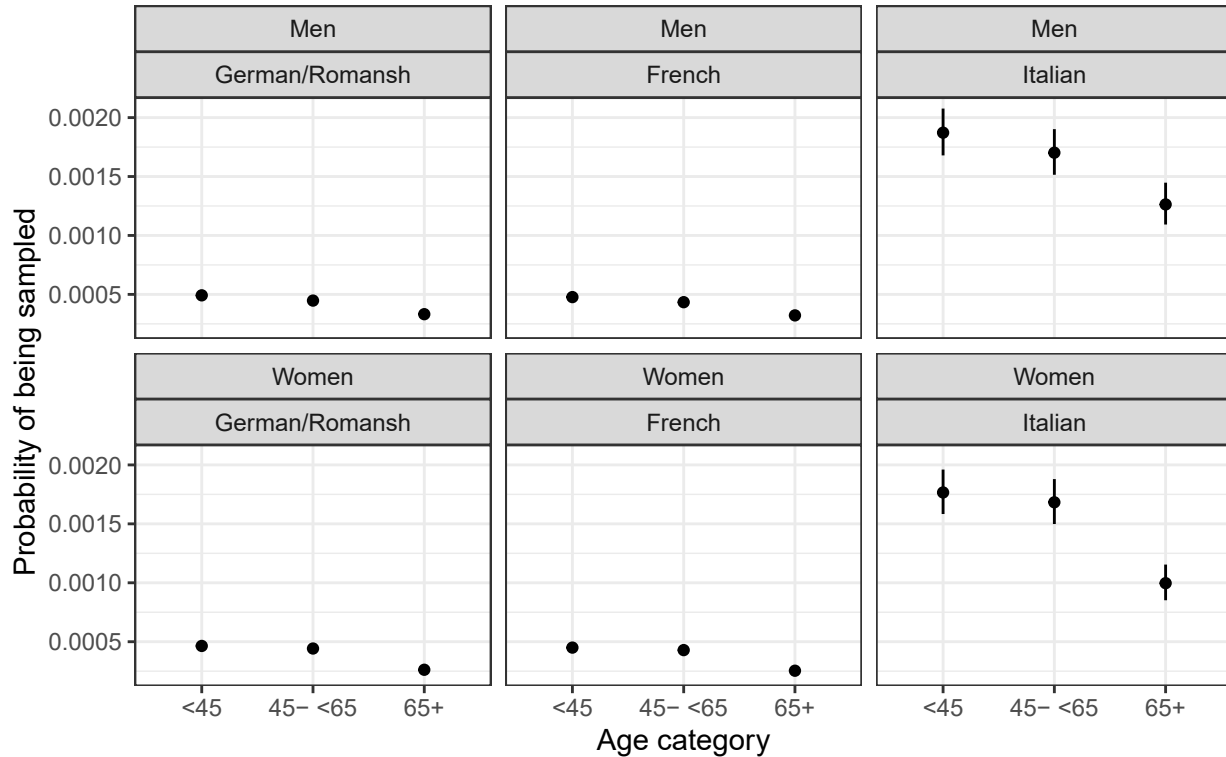
where  $n_u$  is the survey sample size in  $u$ , with additional variables living with partner (*partner*: 0 No, 1 Yes), working situation (*work*: 1 Employed, 2 Unemployed, 3 Retired, 4 Other (e.g. apprenticeship) and highest attained education (*educ*: 1 Compulsory, 2 Secondary, 3 Tertiary). We use noninformative centered Gaussian priors for  $\gamma_z$ ,  $z \leq 5$ , i.e.  $p(\gamma_z | \theta_i, \rho_i^2) = N(0, 1000)$ ,  $z \leq 5$ . We construct nonresponse weights as  $NRW_{iw,u} = 1/(1 - Prob(Z_{iw,u} = 1))$  for  $i \leq n_u, u \in \{1, 2\}, w \geq 2$ . Calibration weights are defined as  $CW_{iw,u} = NRW_{iw,u} \cdot SW_{i,u}$ ,  $i \leq n_u, u \in \{1, 2\}, w \geq 2$ . For Bayesian calculations and model building we used the Integrated Nested Laplace Approximation (INLA) approach (R INLA (<https://www.r-inla.org/>)).

### Construction of sampling weights for survey waves in December 2020 and onwards



agecat	female	lregion	pred_ci
<45	Women	German/Romansh	0.046%, 95% CI (0.043%, 0.05%)
45- <65	Women	German/Romansh	0.044%, 95% CI (0.041%, 0.048%)
65+	Women	German/Romansh	0.026%, 95% CI (0.023%, 0.03%)
<45	Men	German/Romansh	0.049%, 95% CI (0.046%, 0.053%)
45- <65	Men	German/Romansh	0.045%, 95% CI (0.041%, 0.049%)
65+	Men	German/Romansh	0.033%, 95% CI (0.029%, 0.037%)
<45	Women	French	0.045%, 95% CI (0.041%, 0.049%)
45- <65	Women	French	0.043%, 95% CI (0.039%, 0.047%)
65+	Women	French	0.025%, 95% CI (0.022%, 0.029%)
<45	Men	French	0.048%, 95% CI (0.043%, 0.052%)
45- <65	Men	French	0.043%, 95% CI (0.039%, 0.048%)
65+	Men	French	0.032%, 95% CI (0.028%, 0.037%)
<45	Women	Italian	0.177%, 95% CI (0.158%, 0.196%)
45- <65	Women	Italian	0.168%, 95% CI (0.15%, 0.188%)
65+	Women	Italian	0.1%, 95% CI (0.085%, 0.115%)
<45	Men	Italian	0.187%, 95% CI (0.168%, 0.208%)
45- <65	Men	Italian	0.17%, 95% CI (0.152%, 0.19%)
65+	Men	Italian	0.126%, 95% CI (0.109%, 0.145%)

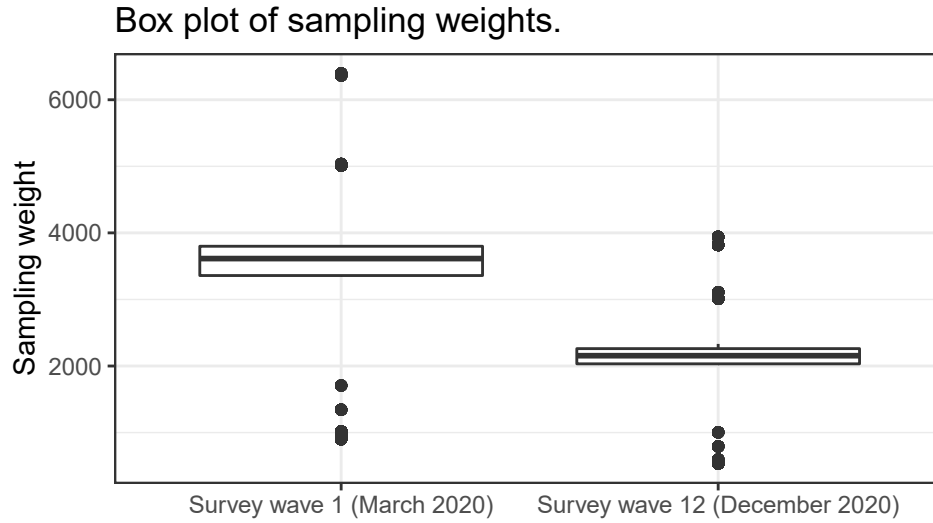
Probability of being sampled from the Swiss population.



Points indicate posterior mean estimates. Lines indicate 95% posterior credible intervals.

A box plot of the sampling weights for the December\* 2020 survey sample and the recalibration sampling weights for the March 2020 survey sample are shown in the next Figure. For the December 2020 survey sample the median of sampling weights is 2155, the 1st quantile 2033, the third quantile 2261, minimum

534, maximum 3939. For the March 2020 survey sample the median of the sampling weights is 3614, the 1st quantile 3359, the third quantile 3799, minimum 900, maximum 6395.



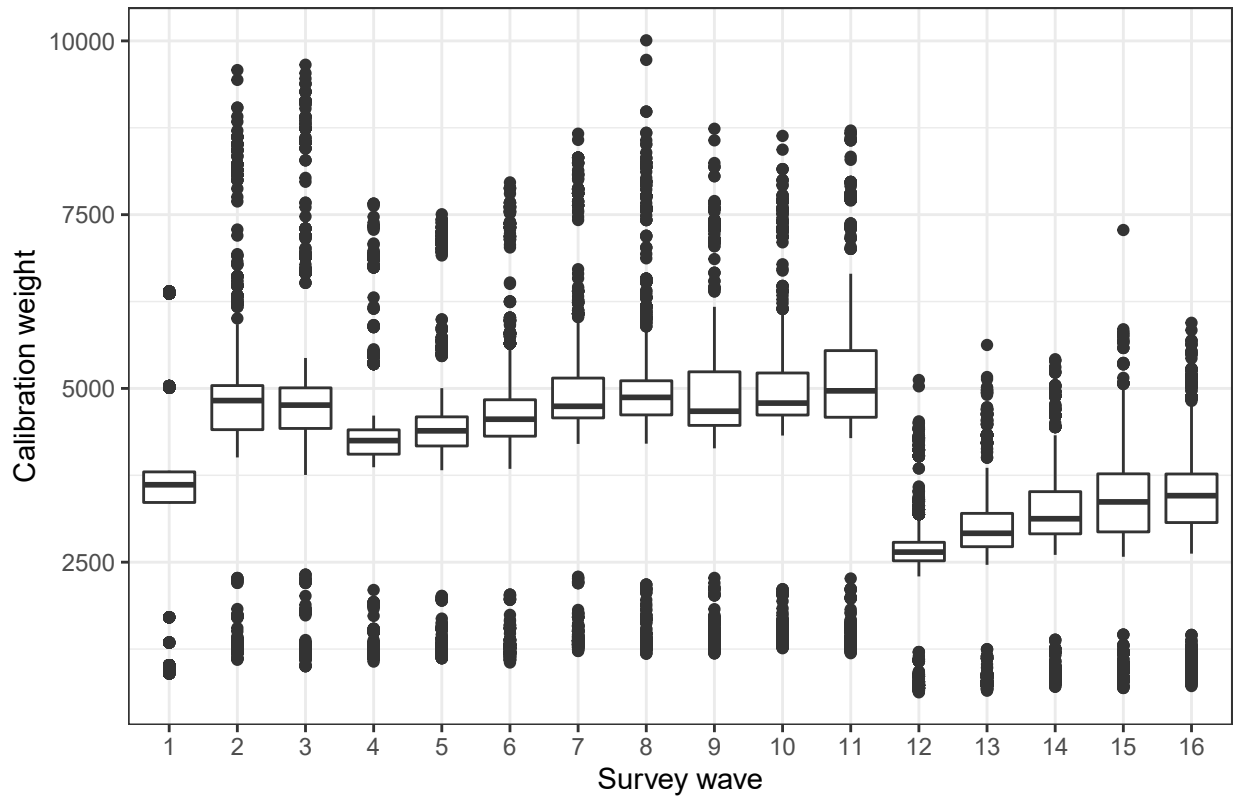
We used the sampling design based on age categories, gender and language region to reconstruct the Swiss population based on the sampling weights. The following table shows the sampling weighted estimated counts with 95% CI (column *Weighted estimate, lci, uci*).

agecat	female	lregion	Weighted estimate	lci	uci	num_pop
<45	Men	German/Romansh	1120183	1043386	1196980	1138490
45- <65	Men	German/Romansh	856388	785786	926990	851030
65+	Men	German/Romansh	551196	487328	615064	535498
<45	Women	German/Romansh	1116290	1036223	1196357	1141840
45- <65	Women	German/Romansh	863702	792359	935045	853534
65+	Women	German/Romansh	557574	481212	633936	537073
<45	Men	French	387760	335017	440503	399229
45- <65	Men	French	269802	223369	316235	275497
65+	Men	French	177099	133747	220451	167685
<45	Women	French	404404	348891	459917	411995
45- <65	Women	French	284504	236667	332341	284307
65+	Women	French	196950	145075	248825	173047
<45	Men	Italian	67818	56472	79164	60468
45- <65	Men	Italian	54004	43401	64607	54279
65+	Men	Italian	32472	22934	42010	39374
<45	Women	Italian	71316	59334	83298	62992
45- <65	Women	Italian	53460	42838	64082	56544
65+	Women	Italian	29087	18798	39376	41017

### Calibration weights

The next figure shows the box plots of calibration weights, by survey wave.

Box plot of calibration weights.



The next table describes the sum of calibration weights for each survey wave. This sum should approximate the underlying 2018 census population of Switzerland aged 15 years or older (N=7.1 million individuals).

wave	calib_weight
1	7088587
2	7100196
3	7093542
4	7097391
5	7099551
6	7096720
7	7086829
8	7095430
9	7090658
10	7095382
11	7091942
12	7098594
13	7097538
14	7095325
15	7088893
16	7094129

### Sensitivity analyses

We perform three sensitivity analyses:

- 1) A model approach with age category, gender and language region as random effects,
- 2) A model approach with age category, gender and language region as fixed effects,
- 3) We compare estimates from the Bayesian approach with a frequentist approach.

We compare WAIC of sensitivity analyses 1) and 2), and the model in the main section. Without loss of generality we use only the December 2020 survey sample to assess model performance.

### Sensitivity analysis 1)

Let  $Y_{iasl}$  denote the binary outcome whether individual  $i \leq 3381$  with age in age category  $a = \{1, 2, 3\}$ , gender  $s = \{1, 2\}$  and living in language region  $l = \{1, 2, 3\}$  was sampled from the Swiss population aged 15 years or older in 2018. We specify a Bayesian hierarchical logistic regression model (M2) as

$$(M2) \quad Y_{iasl} = \beta_0 + agecat_a + gender_s + lregion_l + \epsilon_{iasl}, \quad i \leq 3381, \quad a = \{1, 2, 3\}, \quad s = \{1, 2\}, \quad l = \{1, 2, 3\},$$

where  $\beta_0$  is the overall intercept, and  $agecat_a \sim N(0, \tau_{agecat})$ ,  $gender_s \sim N(0, \tau_{gender})$ ,  $lregion_l \sim N(0, \tau_{lregion})$  are unstructured random effects for variables age category ( $agecat1$ : <45 years,  $agecat2$ : 45- <65 years,  $agecat3$ : 65+ years), gender ( $female$ : 0 Men, 1 Women) and language region ( $lregion$ : 1 German/Romansh, 2 French, 3 Italian). We use a noninformative centered Gaussian prior for  $\beta_0$ , i.e.  $p(\beta_0 | \mu_0, \sigma_0^2) = N(0, 1000)$  and that the prior for the unstructured random effects are centered Gaussian distributed with inverse-Gamma distributed log precision parameters  $\tau_{agecat}, \tau_{gender}, \tau_{lregion} \sim \Gamma^{-1}(1, 5/100000)$ .

### Sensitivity analysis 2)

Let  $Y_i \in \{0, 1\}$  denote the binary outcome whether individual  $i \leq 3381$  is sampled from the Swiss population aged 15 years or older in 2018. We define a non-hierarchical logistic regression model (M3) as

$$(M3) \quad Y_i = \beta_0 + \beta^T agecat * female * lregion + \epsilon_i, \quad i \leq 3381,$$

where  $\beta_0$  is the overall intercept, and  $\beta^T = (\beta_1, \beta_2, \dots, \beta_{17})$  is a vector of estimates from an interaction term (a+b+c+a:b+a:c+b:c) between variables age category ( $agecat1$ : <45 years,  $agecat2$ : 45- <65 years,  $agecat3$ : 65+ years), gender ( $female$ : 0 Men, 1 Women) and language region ( $lregion$ : 1 German/Romansh, 2 French, 3 Italian). We use a Bayesian approach with noninformative centered Gaussian priors for  $\beta_z$ ,  $z \leq 17$ , i.e.  $p(\beta_z | \mu_z, \sigma_z^2) = N(0, 1000)$ ,  $z \leq 17$ .

The following table shows the model performance measured by WAIC for the specified models (M1), (M2) and (M3).

model	WAIC
Main model (M1)	147.9
Sensitivity model (M2)	152.5
Sensitivity model (M3)	151.9

The specified model in the main analysis (M1) showed the lowest WAIC, i.e. best model performance (WAIC=147.9).

### Sensitivity analysis 3)

The following table shows estimates and 95% CI from a frequentist logistic regression model (columns  $est\_glm$ ,  $lci\_glm$ ,  $uci\_glm$ ) and posterior means and 95% credible intervals from the Bayesian model of the specified model in (M3) (columns  $est\_inla$ ,  $lci\_inla$ ,  $uci\_inla$ ).

	est_glm	lci_glm	uci_glm	est_inla	lci_inla	uci_inla
(Intercept)	-7.633	-7.717	-7.550	-7.633	-7.718	-7.550
factor(agecat)2	-0.073	-0.203	0.058	-0.073	-0.204	0.058
factor(agecat)3	-0.348	-0.515	-0.181	-0.348	-0.518	-0.183
female	-0.065	-0.185	0.055	-0.065	-0.185	0.055
factor(lregion)2	-0.043	-0.210	0.123	-0.043	-0.212	0.121
factor(lregion)3	1.468	1.275	1.661	1.469	1.273	1.660
factor(agecat)2:female	0.059	-0.127	0.245	0.059	-0.127	0.245
factor(agecat)3:female	-0.164	-0.413	0.084	-0.164	-0.414	0.084
factor(agecat)2:factor(lregion)2	-0.015	-0.280	0.251	-0.014	-0.282	0.251
factor(agecat)3:factor(lregion)2	0.038	-0.303	0.379	0.038	-0.308	0.375
factor(agecat)2:factor(lregion)3	-0.142	-0.440	0.157	-0.142	-0.442	0.156
factor(agecat)3:factor(lregion)3	-0.354	-0.744	0.036	-0.355	-0.752	0.029
female:factor(lregion)2	0.017	-0.220	0.254	0.017	-0.221	0.254
female:factor(lregion)3	0.016	-0.258	0.290	0.016	-0.259	0.291
factor(agecat)2:female:factor(lregion)2	-0.001	-0.376	0.374	-0.001	-0.377	0.375
factor(agecat)3:female:factor(lregion)2	0.049	-0.448	0.547	0.050	-0.450	0.548
factor(agecat)2:female:factor(lregion)3	-0.073	-0.497	0.351	-0.073	-0.498	0.351
factor(agecat)3:female:factor(lregion)3	-0.174	-0.765	0.416	-0.175	-0.771	0.413

The frequentist model reveals similar estimates as the Bayesian approach in the main analysis.